

Hierarchical models for SBI

- split hidden parameters $\gamma = [Y^G, Y^I]$
 - global: same for an entire population \uparrow
 - individual: different for every instance \uparrow
- } mixed effects model

• example virology Y^G properties of a patient (e.g. disease)

Y^I —||— of individual cells we analyse

epidemiology Y^G properties of a region

Y^I —||— of individual persons

- simulation
 - $Y^G \sim p^S(Y^G)$ once for entire population
 - $Y_i^I \sim p^S(Y^I | Y^G)$ once for every individual / instance

$$X_n = \Phi(Y^G, Y_i^I, \eta_n) \quad \eta_n \sim p^S(\eta | Y^G, Y_i^I) \text{ or } p^S(\eta)$$

$$L = \Phi(Y_i^I, \eta_n) \quad \text{common simplification } \Rightarrow \text{TS}$$

estimation of a non-linear mixed effects model [Arruda et al 2023]

- given $p^s(Y^I | Y^G)$ and $p^s(Y^G)$ as part of simulation
- calculate marginal $p^s(Y^I) = \int p^s(Y^I | Y^G) p^s(Y^G) dY^G$ (either analytically or by learned generative model)
- use simulated TS to train a CNF for individual effects

$$p(Y_n^I | X_n) \approx p^s(Y_n^I | X_n)$$

- determine individual effects for a test set $\{X_n\}_{n=1}^N$ with fixed (unknown) Y^G
 $\{\hat{Y}_{nk}^I \sim p(Y_n^I | X_n)\}_{k=1}^M$ (M posterior samples for Y_n^I for each n)
- find Y^G via numeric optimisation over $p(\{X_n\}_{n=1}^N | Y^G)$ (marginal likelihood)

- under the conditional ($\hat{=}$ given Y^G) iid assumption, marginal likelihood is

$$p^s(\{X_n\}_{n=1}^N | Y^G) = \prod_{n=1}^N \underbrace{p^s(X_n | Y^I)}_{\text{simulation}} \underbrace{p^s(Y^I | Y^G)}_{\text{hierarchical prior for fixed } Y^G} dY^I$$

Bayes rule:
$$p^s(X_n | Y_n^I) = \frac{p^s(X_n) p^s(Y_n^I | X_n)}{p^s(Y_n^I)}$$

$$p^s(\{x_i\}_{i=1}^N | Y_G) = \prod_{i=1}^N p^s(x_i) \cdot \int p^s(Y^I | x_i) \cdot \frac{p^s(Y^I | Y_G)}{p^s(Y^I)} dY^I$$

$$\hat{Y}_G = \operatorname{argmin}_{Y_G} -\log p^s(\{x_i\}_{i=1}^N | Y_G) \quad (\text{standard maximum likelihood estim.})$$

$$= \operatorname{argmin}_{Y_G} \sum_{i=1}^N -\log \left[\mathbb{E}_{Y^I \sim p^s(Y^I | x_i)} \left[\frac{p^s(Y^I | Y_G)}{p^s(Y^I)} \right] \right]$$

approximate by empirical average over $p(Y^I | x_i)$

$$\approx \operatorname{argmin}_{Y_G} \sum_{i=1}^N -\log \left[\frac{1}{M} \sum_{k=1}^M \frac{p^s(Y_{ik}^I | Y_G)}{p^s(Y_{ik}^I)} \right] \quad Y_{ik}^I \sim p(Y^I | x_i)$$

$\underbrace{\hspace{10em}}_{CMF}$

solve by standard non-linear optimization (gradient decent, Newton, BFGS)

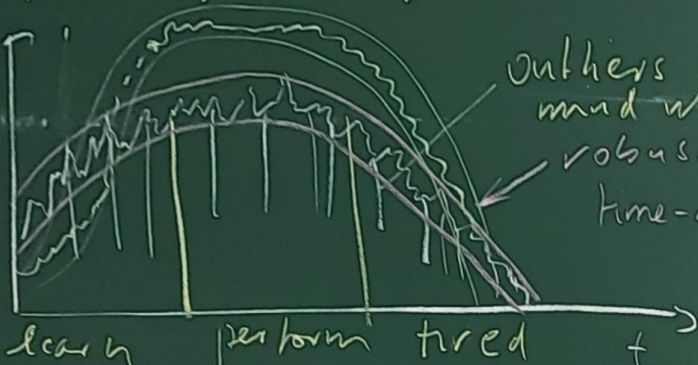
• application: mRNA transfection ($\hat{=}$ how mRNA vaccines enter the cells)

variant: non-stationary time series

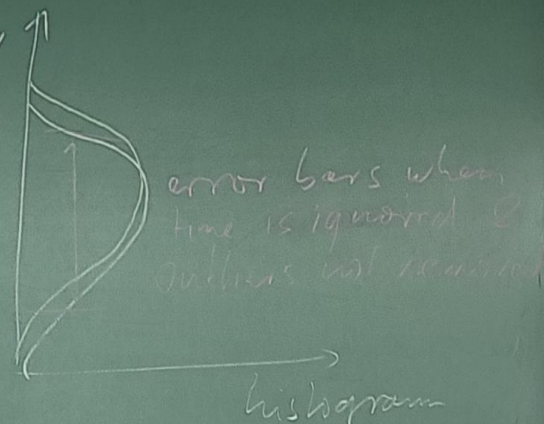
Y_G	parameters	constant over time
Y_t^I	— —	vary over time

eg. psychological experiment.

accuracy
speed



accuracy



eg Y_{it}^I as random walk: $Y_{it}^I = Y_{i,t-1}^I + f(Y_i^G, Y_{i,t-1}^I, \eta_t)$ most information is lost

• brain SBI model with recurrent summary network (eg. LSTM network)

$\mathbb{R}^d \ni h(X_{i,1:t})$ — observations for instance i from timesteps $1 \dots t$

$$\hat{p}_{i,h} = \underset{p_{i,h}}{\operatorname{argmin}} \sum_{i=1}^N \sum_{t=1}^{T_i} -\log p(Y_{it}^I, Y_i^G | h(X_{i,1:t}))$$

cNF
summary network

Y_i^G : stationary parameters of instance i , Y_{it}^I : non-stationary parameters of instance i at time t

Simulation: (b) $Y_i^G \sim p^S(Y_i^G)$ $Y_{i0}^I \sim p^S(Y_{i0}^I | Y_i^G)$

(1) for $t = 1, \dots, T_i$

(a) $Y_{i:t}^I = Y_{i:t-1}^I + f(Y_i^G, Y_{i,t-1}^I, \eta_i^I)$ η_i^I noise for $Y_{i:t}^I$

(b) $X_{i:t} = \Phi(Y_{i:t}^I, \eta_i^X)$ η_i^X noise for $X_{i:t}$

$$\hat{p}(Y_{i:t}^I | Y_i^G | \hat{h}(X_{i:t}^{Obs}))$$

